

# Theory of Stochastic Processes

## 11. Queues

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# Handouts & Announcements

## Handouts:

- Slides (this one)
- Copy of Chapter 8 of PRP
- Copy of §11.1 – 11.3 of PRP

## Announcements:

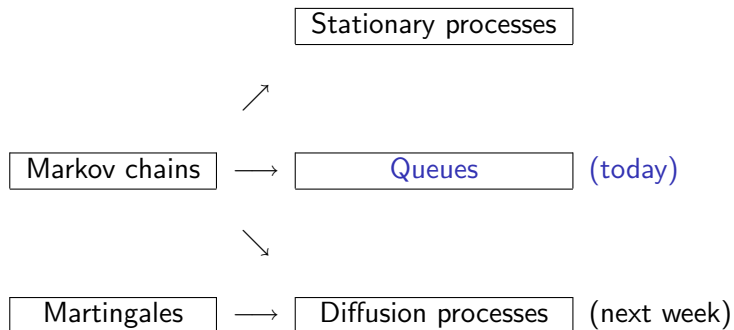
- About the final exam (**important!**)
  - The final exam is held on July 20 (Thu), 10:25–, 90 min.
  - Contact me if you cannot attend it due to some unavoidable reasons.
  - **The exam is open-book and open-note.** Write your answer in English.
  - The exam will cover the whole material up to the next lecture (July 6).  
At least one question will be from the topics before the midterm exam.

# Outline today

- 1 Where are we now?
- 2 Queues
  - Examples
  - M/M/1
  - M/G/1
- 3 Recommended problems

# Where are we now?

Dependence of topics



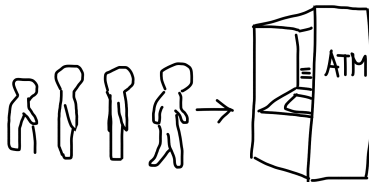
# Queues = waiting lines

Today's topic

## Example

Consider an ATM which serves customers.

- Customers arrive at ATM according to a stochastic process.
- The service time is also random.



# Queueing theory

Other examples of queues:

- security check at Narita airport
- hospitals
- Starbucks coffee
- homework you have to do

The [queueing theory](#) concerns a mathematical study of queueing systems.

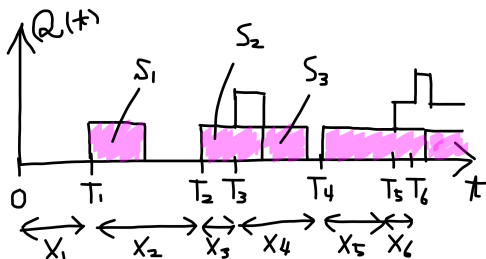
Remark:

- In Japanese, a queue is called “待ち行列”.
- Do not confuse it with “行列” (a matrix).

# Notation

We only consider **single-server** queues.

- $T_n$  is the time when the  $n$ -th customer arrives.
- $X_n = T_n - T_{n-1}$  is the **interarrival time**, where  $T_0 = 0$ .
- $S_n$  is the **service time** of the  $n$ -th customer.
- $Q(t)$  is the number of waiting customers at  $t$ , including ones served.



## Definition

A **queueing model** is a pair of sequences  $\{X_n\}$  and  $\{S_n\}$  of independent random variables.

Queueing models are classified by the distributions of  $X_n$  and  $S_n$ , together with the number of servers ( $= 1$  here). In particular,

- **M/M/1**:  $X_n$  and  $S_n$  are exponential (M stands for **Markov**).
- **M/G/1**:  $X_n$  is exponential and  $S_n$  is any (G stands for **general**).

Remark:

- Remember that interarrival times of a Poisson process are exponential.
- The notation  $A/B/s$  is called Kendall's notation.



# An important quantity

## Definition

The **traffic intensity**  $\rho$  is defined by

$$\rho = \frac{E[S]}{E[X]} = \frac{\text{(mean service time)}}{\text{(mean interarrival time)}}.$$

We may expect that

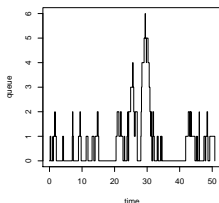
- If  $\rho < 1$ , then  $\{Q(t)\}$  is “stable”.
- If  $\rho > 1$ , then  $\{Q(t)\}$  is “unstable”.

Let us check out this claim for M/M/1 and M/G/1.

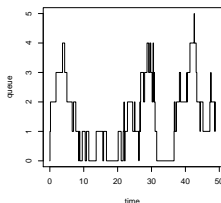
- $M(\lambda)/M(\mu)/1$ :  $X_n$  and  $S_n$  are exponential with parameters  $\lambda$  and  $\mu$ , respectively. The pdfs are  $f_X(x) = \lambda e^{-\lambda x}$  and  $f_S(x) = \mu e^{-\mu x}$ .
- The traffic intensity is

$$\rho = \frac{E[S]}{E[X]} = \frac{1/\mu}{1/\lambda} = \frac{\lambda}{\mu}.$$

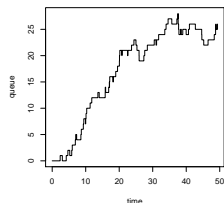
- Sample paths of  $Q(t)$ :



$\rho = 0.5$



$\rho = 1$



$\rho = 1.5$

An R code is available from the course web site.

# Stability of M/M/1

## Theorem 11.2.8

Assume the  $M(\lambda)/M(\mu)/1$  model. Then  $Q(t)$  is a **birth-death process**, that is, a continuous-time Markov chain with the generator

$$\mathbf{G} = \begin{pmatrix} -\lambda & \lambda & & 0 \\ \mu & -(\lambda + \mu) & \lambda & \\ & \mu & -(\lambda + \mu) & \ddots \\ 0 & & \ddots & \ddots \end{pmatrix}$$

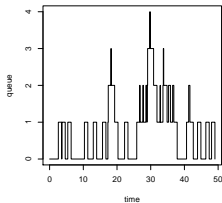
In particular, the stationary distribution exists if and only if  $\rho = \lambda/\mu < 1$ .

## Review exercise

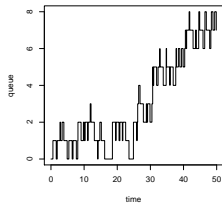
Let  $\rho < 1$ . Show that the stationary distribution is  $\pi_k = (1 - \rho)\rho^k$  and the mean number of waiting customers is  $\rho/(1 - \rho)$ .

Now let us go on to the M/G/1 model.

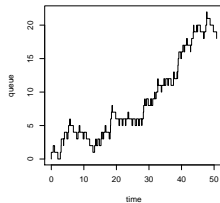
- Sample paths of  $Q(t)$  when  $S_n \equiv 1$  (deterministic).



$$\rho = 0.5$$



$$\rho = 1$$



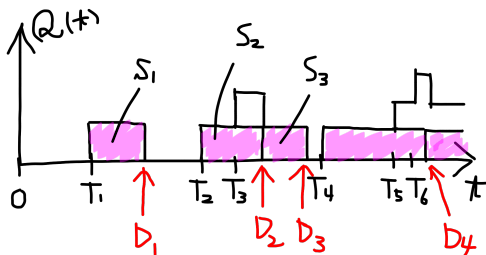
$$\rho = 1.5$$

## Exercise

$Q(t)$  is **no longer** a Markov chain. Why?

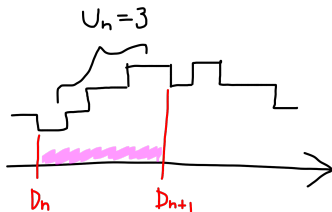
# The idea

- Let  $D_n$  be the **leaving time** of the  $n$ -th customer. In other words,  $D_n$  is the  $n$ -th decreasing point of  $Q(t)$ .
- We will see that  $Q(D_n)$  is a discrete-time Markov chain, where  $Q(D_n)$  is interpreted as  $Q(D_n+)$ .

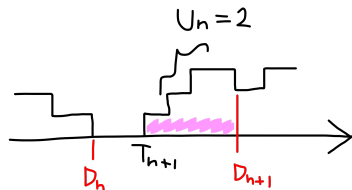


# Find a Markov chain

- Let  $U_n$  be the number of arrived customers during the service time of the  $(n + 1)$ -th customer.
- $U_n$  is independent of  $\{Q(t)\}_{t \leq D_n}$  because of the Markov property of arriving times.



$$Q(D_n) > 0$$



$$Q(D_n) = 0$$

## Exercise

Check that  $Q(D_{n+1}) = U_n + \max(Q(D_n) - 1, 0)$ .

### Theorem 11.3.4

Assume the  $M(\lambda)/G/1$  model. Then the discrete-time process  $\{Q(D_n)\}$  is a Markov chain with the transition matrix

$$\mathbf{P} = \begin{pmatrix} \delta_0 & \delta_1 & \delta_2 & \cdots \\ \delta_0 & \delta_1 & \delta_2 & \cdots \\ 0 & \delta_0 & \delta_1 & \cdots \\ 0 & 0 & \delta_0 & \cdots \\ \vdots & \vdots & \vdots & \ddots \end{pmatrix}, \quad \delta_j = P(U_n = j) = E \left[ \frac{(\lambda S_n)^j}{j!} e^{-\lambda S_n} \right].$$

### Theorem 11.3.5

The unique stationary distribution  $\pi$  exists if and only if  $\rho < 1$ . In that case, the generating function of  $\pi$  is

$$G(s) = \sum_j \pi_j s^j = (1 - \rho)(s - 1) \frac{M_S(\lambda(s - 1))}{s - M_S(\lambda(s - 1))},$$

where  $M_S(\theta) = E[e^{\theta S_n}]$ .

See §11.3 for proofs and further details.

# Recommended problems

## Recommended problems:

- §8.4, Problems 1, 2, 3, 4\*, 5\*.
- §11.2, Problems 3, 7\*.
- §11.3, Problem 1\*.

The asterisk (\*) shows difficulty.